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The hammer and the scalpel: On the economics of indiscriminate versus targeted isolation policies during pandemics [☆]

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ABSTRACT

We develop a simple dynamic economic model of epidemic transmission designed to be consistent with widely used biological models of the transmission of epidemics, while incorporating economic benefits and costs as well. Our main finding is that if the technology for tracking infected individuals is sufficiently good, targeted testing and isolation policies deliver large welfare gains relative to optimal policies when these tools are not available. Much of this welfare gain comes from isolating infected individuals rather than testing them. When the tracking technology is not very good, the gains from targeted testing and isolation are small. The message of our analysis is that the returns to improving tracking technologies are very large.

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1. Introduction

Pandemics force households, firms, and governments to make cruel choices among unhappy alternatives. Much economic activity is enhanced by close person-to-person contact. Unfortunately, this kind of contact typically allows viruses to be more easily transmitted from person to person. Much of the economic literature on epidemics studies the trade-offs between the losses to economic activity associated with limiting contact and the gains from reduced transmission of the virus, including reduced healthcare costs, lower strains on hospitals, and fewer deaths. The response to the coronavirus epidemic in most Western countries has been to limit contacts by limiting economic activity. Some countries—most notably, South Korea, Singapore, Hong Kong, and Taiwan—have limited economic activity to a lesser extent and have supplemented the modest limitations with aggressive policies of targeted testing, contact tracing, and isolation.

In this paper, we develop a version of a fairly standard macroeconomic model of epidemics, incorporate testing and isolation policies into it, and ask to what extent testing policies of targeted testing and isolation can achieve better outcomes. In the version of our model calibrated to the tracking technology available in South Korea, we find that a policy of targeted

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testing and isolation yields substantial welfare gains. If testing and isolation policies are optimally designed, economic activity must be curtailed to a much more limited extent, and the number of deaths is substantially smaller than if testing and isolation policies are not available. We argue that relative to no testing, untargeted testing without using a tracking technology yields only modest benefits. We also argue that even if testing resources are not available, targeted isolation without testing yields about two-thirds of the welfare gains from a targeted testing and isolation policy.

We then calibrate a version of our model to the tracking technologies available in Germany and Australia and find that targeted testing and isolation policies yield at best modest benefits relative to a policy of untargeted testing. The main reason for the difference is that the fraction of infected people who are successfully tracked is much smaller in our Germany and Australia experiments than in our South Korea experiment. Our analysis implies that the returns to tracking, contact tracing, and the like are very large.

The simple dynamic economic model of epidemic transmission developed here is designed to be consistent with widely used SIR (Susceptible, Infected, Recovered) biological models of the transmission of epidemics (see Atkeson (2020b) for a primer), while incorporating economic benefits and costs as well. We choose a formulation that makes it possible to analyze the benefits and costs of various policies. The main stand we take is that social proximity has benefits by allowing for economic activity to take place. We have in mind that certain types of production activities require groups of people to work in close proximity to one other. The obvious example of such an activity is assembly line production. In other activities, individuals derive value from social proximity in consumption. Examples of such activities are watching live performances of plays or rock concerts. While substitutes are available for production and consumption with social proximity (artisanal production as opposed to assembly line production, or televised rock concerts versus live ones), it is often the case that it is cheaper to provide a given good or service with high levels of contact than with low levels of contact.¹ Social proximity has costs when such proximity allows viruses to be relatively easily transmitted.

The standard SIR model in epidemiology has three types of agents: susceptible (or not yet infected) agents, infected agents, and recovered agents (who may be alive or dead). To allow for testing, we extend the model to allow for two types of infected agents: those known to be infected and those not known to be infected. In our economic model, agents engage in a variety of economic activities. Each economic activity is associated with a given number of “meetings” with other agents. These activities are combined to produce a final output good. The virus is transmitted with an exogenous activity-specific probability in a meeting between an infected and susceptible agent. We assume that activities with low transmission probabilities also have low economic value.

The planner seeks to maximize the present discounted utility of consumption net of costs of treating infected agents and of death costs. In our model, absent any testing, the planner excludes from any activity agents who are known to be infected and allocates some of the agents whose types are not known to the activity with the lowest probability of transmission. Since the lowest probability of transmission activity has the lowest economic value, this policy tends to reduce output, but it saves lives. The quantitative version of our model generates output declines and death reductions broadly similar to those in Eichenbaum et al. (2020b) and Glover et al. (2020).

We measure the welfare gains from optimal policy as the permanent percentage increase in consumption that would give the planner the same utility as under no policy. For reference, we note that the loss in welfare in the no-intervention economy relative to the no-pandemic economy is 6.66%. We show that the optimal policy yields a welfare gain of roughly 0.6% relative to no intervention. We then introduce a costly testing technology.² The planner can choose to test a fraction of the population whose types are not known. We assume the test perfectly reveals whether an agent is infected. We show that optimal policy with this type of untargeted testing yields a welfare gain of 0.7% relative to no intervention. That is, untargeted testing delivers gains of only 0.1% relative to welfare under optimal policy with no testing.

We allow for targeted testing by assuming that each agent whose type is not known is associated with a signal that he is infected. We think of this signal as combining information from a variety of sources. One example is contact tracing, which involves tracing people an infected person has come into contact with, persons whom these contacts contacted, and so on. This signal is informative in the sense that the probability of receiving the signal is higher for infected agents than for susceptible agents. We assume the signal is not perfectly revealing in that the probability an infected person receives the signal is strictly less than one. Out of these agents with the signal, the planner chooses the fraction to test. We choose the signal probabilities to be consistent with data from South Korea. That data suggest that 38% of infected people and 0.44% of susceptible people are associated with the signal. We show that the welfare gain from optimal policy with targeted testing relative to no intervention is roughly 3% of consumption forever. That is, targeted testing allows for a dramatic gain in welfare relative to a no-intervention policy.

In our model, targeted testing allows the planner to very precisely target some agents in order to isolate them. We separate out the effects of isolation from the informational gains to testing by considering a version of the model in which the planner cannot test after receiving the signal.³ The planner simply chooses the fraction of individuals with a signal to

¹ Obviously across goods those with typically low levels of contact may provide greater economic value than those goods with high levels of contact.

² Clearly, if testing is costless then it is optimal to test everyone in the population. We find more interesting and address below the case in which testing is so costly that it is cheaper to simply isolate the individuals who are suspected of being infected rather than to test.

³ Testing can, of course, be very valuable in learning about the current state of the system and the parameters governing its evolution. Our focus here is on testing to isolate infected individuals, rather than testing as a learning device.

isolate. We show that under optimal isolation, the welfare gains are 2% of consumption forever. That is, roughly two-thirds of the gains from optimal testing can be realized by forgoing testing and simply isolating agents suspected of being infected.

We then consider a calibration of our tracing technology that is intended to be consistent with the German and Australian experience. It turns out that in these countries the tracing technology was substantially less effective than in South Korea. Viewed through the lens of our model, the fraction of infected people who emit the signal is much smaller in Germany or Australia than it is in South Korea. We find that calibrated to the German or Australian experience, testing and isolation strategies are only marginally effective at improving outcomes relative to optimal policy without these strategies.

We also conduct a variety of sensitivity exercises by varying the probability of receiving the signal. We show that if 60% of infected agents and 3% of susceptible agents are associated with the signal, the welfare gains relative to no intervention are about 5.5%. That is, the welfare loss from the pandemic is only about 1% of consumption. We also show that if a relatively small fraction of infected agents receives the signal, then the welfare gains are also smaller.

Our findings with respect to South Korea, Germany, Australia and our sensitivity analyses make clear that the returns to improving tracing technologies are extremely large. In this sense, our findings suggest that the policies advocated by Romer and Garber (2020, March), Romer (2020), and Holtemöller (2020) are most useful when supplemented with aggressive and accurate contact tracing.

1.1. Relation to other recent papers

Here we present a discrete-time version of the standard continuous-time SIR epidemiology model outlined (for economists) by Atkeson (2020b). A long (and growing) list of papers emphasizes the trade-offs between the losses from restrictions on economic activity and the losses from allowing the virus to spread. See, among many others, Alvarez et al. (2020), Atkeson (2020b), Atkeson (2020a), Azzimonti et al. (2020), Baqaee et al. (2020), Bodenstein et al. (2020), Eichenbaum et al. (2020b), Farboodi et al. (2020), Garriga et al. (2020), Guerrieri et al. (2020), Hall et al. (2020), Jones et al. (2020), Kaplan et al. (2020), Krueger et al. (2020), Moser and Yared (2020), Rampini (2020), and Toxvaerd (2020). While none of these papers focuses on the role of testing, some other recent papers do. See, for example, Eichenbaum et al. (2020a), Berger et al. (2020), Acemoglu et al. (2020), and Piguillem et al. (2020). Our findings complement the results in this literature regarding the desirability of testing. In particular, Eichenbaum et al. (2020a) and Piguillem et al. (2020) also emphasize the critical role of isolation.

2. An economic model with contagion

Consider a discrete time infinite horizon model with a continuum of individuals on the unit interval. At any date, an individual is either *susceptible* S , *infected* I (and thus contagious), or *recovered* R , (where *recovered* can be either dead or alive). Of the infected, $I - \tilde{I}$ agents are known to be infected, and \tilde{I} agents are not known to be infected (hereafter *unknown infected*).

For the bulk of our analysis, we assume that the planner knows which agents are recovered (R_t) and which are known to be infected ($I_t - \tilde{I}_t$) but cannot tell apart susceptible (S_t) and unknown infected agents (\tilde{I}_t). That is, the planner must treat susceptible and unknown infected agents in the same way but can treat other agents in different ways. The state of the system for the planner is $(S_t, I_t, \tilde{I}_t, R_t)$, where the variables denote the fraction of agents who are susceptible, infected, unknown infected, and recovered (dead or alive) at date t . The implicit assumption here is that costless tests are available to tell whether a person has been infected in the past.

We relax this assumption by also considering a version of the model in which these tests are not available. In this *five state* version of the model, recovered agents are partitioned into two types, those who are recovered after not having been known to be infected (\tilde{R}_t) and those who are recovered after having been known to be infected ($R_t - \tilde{R}_t$). In this version, the state of the system is $(S_t, I_t, \tilde{I}_t, R_t, \tilde{R}_t)$. The planner treats those who are recovered after not having known to be infected (\tilde{R}_t) in the same way as susceptible people. Our benchmark model is the four state version, rather than the five state version, primarily because the computational time for the five state model is orders of magnitude greater than the four state model. It turns out that the four state model is a good approximation of the five state model for our baseline experiments. This finding leads us to be confident that the four state model is a good approximation for the other experiments as well.

An infected person dies with probability $\gamma\delta$, stays alive with probability $\gamma(1 - \delta)$, and stays infected with probability $(1 - \gamma)$. We assume, as is conventional in much of the literature, that infection confers permanent immunity so that a recovered person always stays recovered—again, either dead or alive. If $S_0 + I_0 = 1$ (all individuals start as susceptible or infected), this assumption ensures that for all t , $(1 - \delta)R_t$ fraction of initially alive people are alive and δR_t fraction of initially alive people are dead.

In our model, economic activity is associated with meetings or interactions. Our economy has N types of intermediate goods labeled $i \in \{1, \dots, N\}$. Good i is produced using activity i . Each activity of type i requires M_i meetings for each person engaged in that activity. The technology for producing good i is given by

$$y_{it} = b_i \tilde{I}_{it}, \quad (1)$$

where y_{it} is the amount of good i , b_i is a technology parameter that depends on activity i and the number of meetings, and \tilde{L}_{it} is the amount of effective labor allocated to good i . Type S and type R agents each supply one unit of effective labor, and infected agents supply ξ units of effective labor. The intermediate goods are combined into a final consumption good by a CES aggregator, so that the amount of the final good Y_t produced in period t is

$$Y_t = \sum_i (y_{it}^{\frac{\sigma-1}{\sigma}})^{\frac{\sigma}{\sigma-1}}. \quad (2)$$

In the presence of a pandemic, meetings also lead to type-specific transmission of the virus. The purpose of indexing meetings by type i and allowing the transmission rate to depend on meeting type i is to allow the framework to consider multiple types of policy interventions, such as prohibiting or decreasing particular types of meetings. For instance, the probability of transmission while chatting on a sidewalk can depend on whether both people are wearing masks. Meeting while wearing masks can be considered a different type of meeting than meeting while not wearing masks. Meetings that occur only after each participant has had his temperature checked can be considered a different type of meeting from meetings where such temperature checks do not occur. Further, we later consider the costs of various policy interventions. While the rate of transmission for two workers standing next to each other might be the same regardless of what they are producing, the cost to society of reducing such meetings may very well depend on whether they are producing ventilators or academic papers (since the latter can be more easily moved online). In such a case, these two activities would be considered different types of meetings.

2.1. Model with no testing

To make the exposition easier, we begin by considering a version of the model with no testing and then introduce testing. In our model, economic activity induces infections and thereby the resulting laws of motion for the state variables. Let L_{it} denote the mass of people assigned to activity i , λ_{it} denote the fraction of agents whose types are unknown (these consist of S and \tilde{I} type agents) assigned to activity i , and μ_{it} the fraction of recovered agents who are assigned to activity i . We assume that agents who are known to be infected are assigned to activity N . Each activity i is characterized by the probability that a susceptible person who meets an infected person gets infected, p_i . Within an activity, meetings are independently drawn. Since the mass of people assigned to activity i is L_{it} , and the mass of infected people assigned to activity i is $\lambda_{it}\tilde{I}_t$, a susceptible person in a single meeting meets an infected person with probability $\lambda_{it}\tilde{I}_t/L_{it}$ and gets infected with probability $p_i\lambda_{it}\tilde{I}_t/L_{it}$. The probability of being infected in M_i meetings is then

$$1 - \left(1 - p_i \frac{\lambda_{it}\tilde{I}_t}{L_{it}}\right)^{M_i}, \quad (3)$$

where

$$L_{it} = \lambda_{it}(S_t + \tilde{I}_t) + \mu_{it}(1 - \delta_t)R_t. \quad (4)$$

Since the mass of susceptible people assigned to activity i is $\lambda_{it}S_t$, the law of motion for the mass of susceptible people in the population is

$$S_{t+1} = S_t - \sum_i \lambda_{it}S_t \left(1 - \left(1 - p_i \frac{\lambda_{it}\tilde{I}_t}{L_{it}}\right)^{M_i}\right). \quad (5)$$

Taking a Taylor series expansion of this law around $\tilde{I}_t = 0$, we obtain a law of motion similar to that in the SIR models, given by

$$S_{t+1} = \left[1 - \sum_i \left(\lambda_{it}\pi_i \frac{\lambda_{it}\tilde{I}_t}{L_{it}}\right)\right] S_t, \quad (6)$$

where $\pi_i = p_i M_i$. We use this approximation in our quantitative assessment of policies.

In terms of the law of motion for \tilde{I} , we assume that a person of type \tilde{I} becomes known to be infected with exogenous probability $\tilde{\tau}$. Since an infected person stays infected with probability $1 - \gamma$, the law of motion for \tilde{I} is

$$\tilde{I}_{t+1} = (1 - \gamma)(1 - \tilde{\tau})\tilde{I}_t + S_t \sum_i \left(\lambda_{it}\pi_i \frac{\lambda_{it}\tilde{I}_t}{L_{it}}\right), \quad (7)$$

and the law of motion for R is

$$R_{t+1} = R_t + \gamma(1 - S_t - R_t). \quad (8)$$

Together with the adding up constraint that is $S_t + I_t + R_t = 1$, the system of equations, (5), (7), and (8) describe the dynamics of the system. This dynamical system has a continuum of steady states, indexed by S , the steady state fraction of susceptible individuals, with R , the steady state fraction of recovered individuals equal to $1 - S$ (and thus $I = 0$).

Given an initial state of the system (\tilde{I}_0, R_0, S_0) and the policy variables λ_{it} and μ_{it} , the approximate dynamical system, with (5) replaced by (6), is identical to the familiar SIR models from epidemiology. A key variable in these models is $\mathcal{R}_{0,t}$ – the number of new infections per susceptible person per infected person, multiplied by the mean number of periods an infected person is infected. From equation (6), using the observation that the mean number of periods infected is $1/\gamma$, we have

$$\mathcal{R}_{0t} \approx \frac{1}{\gamma} S_t \frac{\sum_i \left(\lambda_{it} \pi_i \frac{\lambda_{it} \tilde{I}_t}{L_{it}} \right)}{I_t S_t}. \quad (9)$$

2.2. Model with testing

Next, we introduce testing into the model. At the beginning of each period, each person whose type is not known is associated with a public signal that he is infected.⁴ Let θ_X , $X \in \{S, \tilde{I}\}$ denote the probability this signal is received regarding an individual of type X .⁵ That is, the signal is useful information to the extent that $\theta_I > \theta_S$. With this formulation, the mass of agents who are associated with a signal of infection is $\theta_S S + \theta_I I$. The planner chooses to test the fraction τ_t of these individuals. The test perfectly reveals whether an agent is infected. We assume that the externalities from infection are larger than the value of lost production due to isolation so that the planner optimally isolates all those who test positive.⁶ Thus, the law of motion for \tilde{I} is now

$$\tilde{I}_{t+1} = (1 - \tau_t \theta_I) \left[(1 - \gamma)(1 - \tilde{\tau}) \tilde{I}_t + S_t \sum_i \left(\lambda_{it} \pi_i \frac{\lambda_{it} \tilde{I}_t}{L_{it}} \right) \right], \quad (10)$$

where L_{it} is given by

$$L_{it} = \lambda_{it} \left(S_t + (1 - \tau_t \theta_I) \tilde{I}_t \right) + \mu_{it} (1 - \delta_t) R_t, \quad (11)$$

and output in activity i is given by

$$y_{it} = b_i \tilde{I}_{it} = b_i \left(\lambda_{it} (S_t + (1 - \tau_t \theta_I) \tilde{I}_t) + \mu_{it} (1 - \delta_t) R_t \right). \quad (12)$$

The cost of testing is $C(\tau_t(\theta_S S_t + \theta_I I_t))$, where $C(\cdot)$ is the testing cost function. The role of testing is to remove some of the infected agents whose type is not known from current and future economic activity until they recover. This policy of removing some of the infected agents prevents them from infecting others, in both current and future periods. In addition to the testing cost, removing these agents is costly since they cannot engage in useful economic activity.

Aggregate consumption is given by $Y_t - C(\tau_t(\theta_S S_t + \theta_I I_t))$. The planner's preferences over consumption are given by $\sum_{t=0}^{\infty} \beta^t U(Y_t - C(\tau_t(\theta_S S_t + \theta_I I_t)))$. In addition, infection is associated with a utility cost $Z I_t$, where Z denotes the healthcare and related costs of being infected, and deaths are associated with a utility cost $D \gamma \delta I_t$, where we note that the mass of agents who die in period t is $\gamma \delta I_t$ and the parameter D measures the cost of a life.

The planning problem is then to choose a testing policy τ_t and labor allocation policies λ_{it} and μ_{it} to solve

$$\max \sum_{t=0}^{\infty} \beta^t [U(Y_t - C(\tau_t(\theta_S S_t + \theta_I I_t))) - Z_t(I_t) - D_t(\gamma \delta I_t)],$$

subject to (1), (2), (4), (6), (8), and (10), given the initial conditions $(S_0, I_0, \tilde{I}_0, R_0)$. We will refer to the version of the model with $\theta_I = \theta_S = 1$ as the model with *untargeted* testing and the model with $\theta_I > \theta_S$ as the model with *targeted* testing. Note that the programming problem for the model without testing is simply this programming problem with $\tau_t = 0$ for all t .

In order to understand how policy can be used to affect the course of the infection, consider a simple version of the model with two activities, *work* and *home*. Suppose that work produces higher output but is also associated with higher infection than the home activity, so that $b_{work} > b_{home} = 0$ and $\pi_{work} > \pi_{home} = 0$. Without any testing, it is optimal in general to assign some agents whose types are not known to stay home. This policy reduces economic activity but also reduces

⁴ Since all information is public, there is no difference between what individuals know and what the planner knows. It would be interesting, but beyond the scope of this paper, to consider situations in which information is asymmetric and there is a conflict of interest between private incentives to disclose information and social needs to acquire the information.

⁵ In the five state model we assume that both susceptible people and people who have recovered without being previously publicly known as infected (\tilde{R}_t) emit the signal with probability θ_S . This assumption maintains the informational symmetry between these two types of people.

⁶ This assumption holds in our quantitative analysis.

the virus's rate of transmission. We refer to this policy as *indiscriminate* isolation, since the policy does not discriminate between infected and susceptible agents. Consider next the role of targeted testing in this simple model. Targeted testing allows the planner to isolate some of the infected agents by requiring them to stay at home. In this sense, targeted testing allows for a form of targeted isolation.

2.3. Model with isolation and no testing

As was the case before, agents receive signals at the beginning of the period. In this version of the model, however, the planner does not test but simply isolates some fraction of the agents who have received signals in the current period. In particular, we assume that the planner does not use past signals in isolating individuals. This assumption implies that welfare under the policy is a lower bound for a more elaborate policy that uses the entire history of past signals. The law of motion for \tilde{I} is now

$$\tilde{I}_{t+1} = (1 - \gamma)(1 - \tilde{\tau})\tilde{I}_t + (1 - \tau_t\theta_S)S_t \sum_i \left(\lambda_{it}\pi_i \frac{\lambda_{it}(1 - \tau_t\theta_I)\tilde{I}_t}{L_{it}} \right), \quad (13)$$

where L_{it} is now given by

$$L_{it} = \lambda_{it} \left((1 - \tau_t\theta_S)S_t + (1 - \tau_t\theta_I)\tilde{I}_t \right) + \mu_{it}(1 - \delta_t)R_t, \quad (14)$$

and output in activity i is given by

$$y_{it} = b_i M_i \left(\lambda_{it} \left((1 - \tau_t\theta_S)S_t + (1 - \tau_t\theta_I)\tilde{I}_t \right) + \mu_{it}(1 - \delta_t)R_t \right). \quad (15)$$

Note that, as in the case with testing, the role of isolation is to remove some fraction of the infected people from economic activities where these people may infect others. One advantage of this policy is that it does not require the use of testing resources. A disadvantage of this policy is that some susceptible people are also removed from productive economic activities. Throughout our analysis of isolation policies, we will assume that $\theta_I > \theta_S$.

Since it is always possible to set $\tau_t = 0$ with either testing or isolation, it immediately follows that welfare is higher than it is when these instruments are not available. In our next result, we show that isolation and testing are policies that are typically used in their entirety before the planner resorts to reducing labor allocation in productive economic activities with high rates of infection. In order to understand this result, it is useful to consider a special case with only two activities: work and home. Suppose that $b_{work} = 1$, $b_{home} = 0$, $\pi_{work} > 0$, and $\pi_{home} = 0$. In this case, absent a pandemic, all agents would be assigned to the work activity. With a pandemic and no testing or isolation, it can be optimal to assign some of the workers of unknown type to stay at home. The trade-off of this policy is that assigning a larger number of agents to stay home reduces output but also reduces the rate at which the virus spreads. If only isolation policies are available, then it is possible to show that it is optimal to exhaust all isolation possibilities before assigning any workers of unknown type to stay home. We formalize this result in the following proposition.

Proposition 1. *Suppose that $\pi_{home} = 0$, $b_{home} = 0$, and $\xi = 1$. Then, if $S > \tilde{I}$, $\lambda_{work} < 1$ only if $\tau = 1$.*

The proof of this result is in the appendix. This result also implies that if the cost of testing is sufficiently small, it is optimal to exhaust all testing possibilities. By continuity, it follows also that if π_{home} and b_{home} are not too different from zero, the same result applies.

In summary, these results imply that in general, welfare under a regime with testing or isolation will be higher than welfare under a regime with no testing or no isolation.

3. Dynamics and the effects of social distancing

In this section, in order to obtain intuition about the trade-offs that optimal policy must confront, we illustrate the behavior of our dynamical system for some simple cases. Suppose now that the fraction of agents who are known to be infected relative to the mass of infected agents is constant. Let $q = 1 - \tilde{I}/I$ denote this fraction. Suppose also that $\lambda_{it} = \mu_{it} = 1/N$ for all t . Then, it is easy to show that the law of motion for the mass of infected agents is given by

$$I_{t+1} = I_t(1 - \gamma) + S_t \left(1 - \frac{1}{N} \sum_i \left(1 - p_i \frac{(1 - q)I_t}{S_t + (1 - q)I_t + R_t} \right)^{M_i} \right), \quad (16)$$

and that for the mass of susceptible people is given by

$$S_{t+1} = S_t - S_t \left(1 - \frac{1}{N} \sum_i \left(1 - p_i \frac{(1 - q)I_t}{S_t + (1 - q)I_t + R_t} \right)^{M_i} \right). \quad (17)$$

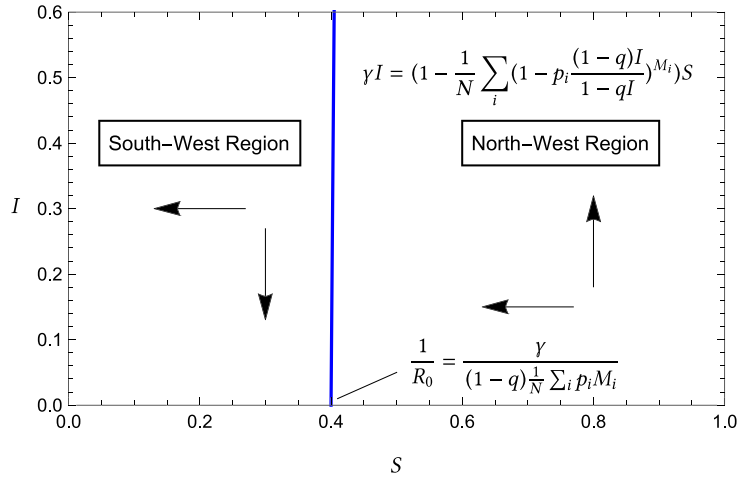


Fig. 1. (S, I) Phase Diagram.

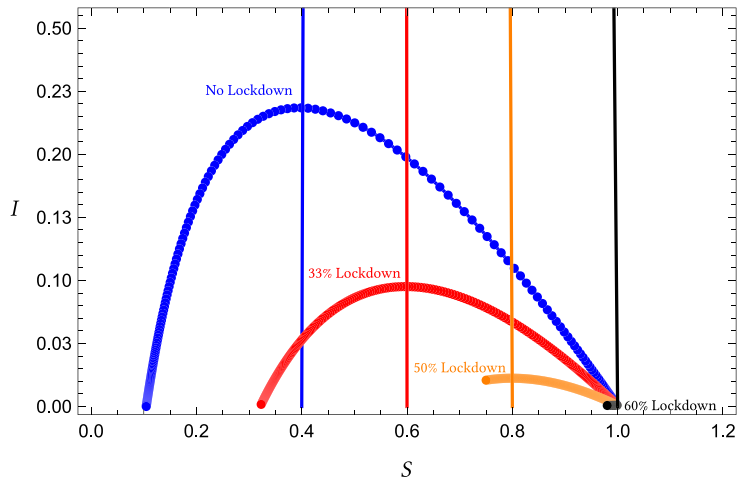


Fig. 2. The Effects of Constant Economic Lockdowns for One Year.

In Fig. 1, we display the evolution of this dynamical system in state space form, with I_t on the y-axis and S_t on the x-axis. Note that equation (17) implies that if $I_t > 0$, $S_{t+1} - S_t < 0$. That is, S moves to the left, or west, in Fig. 1. To see how I_t evolves, we partition (S, I) space into those points to the right and left of the upsloping locus of (S, I) points such that I is constant. We derive this locus by setting $I_{t+1} = I_t = I$, $S_t = S$, delivering

$$\gamma I = \left(1 - \frac{1}{N} \sum_i \left(1 - p_i \frac{(1-q)I}{1-ql}\right)^{M_i}\right) S. \tag{18}$$

Note that this locus intersects the horizontal axis at

$$\lim_{I \rightarrow 0} \frac{\gamma I}{\left(1 - \frac{1}{N} \sum_i \left(1 - p_i \frac{(1-q)I}{1-ql}\right)^{M_i}\right)} = \frac{\gamma}{(1-q) \frac{1}{N} \sum_i p_i M_i} = \frac{1}{\mathcal{R}_0},$$

as indicated in the figure. To the right of this locus, the dynamics of the system are north-west. That is, $S_{t+1} < S_t$ from (17) and $I > 0$, and $I_{t+1} > I_t$ from S being greater than that associated with I being constant (and I_{t+1} being an increasing function of S_t in (16)). To the left of this locus, the dynamics of the system are south-west. Here, again, $S_{t+1} < S_t$ when $I_t > 0$, and $I_{t+1} < I_t$ from S being less than that associated with I being constant.

This implies that if the initial (S_0, I_0) is to the left of the locus, (S_t, I_t) converges to a steady state on the horizontal axis following a south-west path. If the initial (S_0, I_0) is to the right of the locus, (S_t, I_t) converges to a steady state on the horizontal following an arc pattern, with I_t increasing as S_t decreases until (S_t, I_t) crosses the locus, converging again to a steady state on the horizontal axis. Note that this implies all steady states reachable from an initial state with $I_0 > 0$

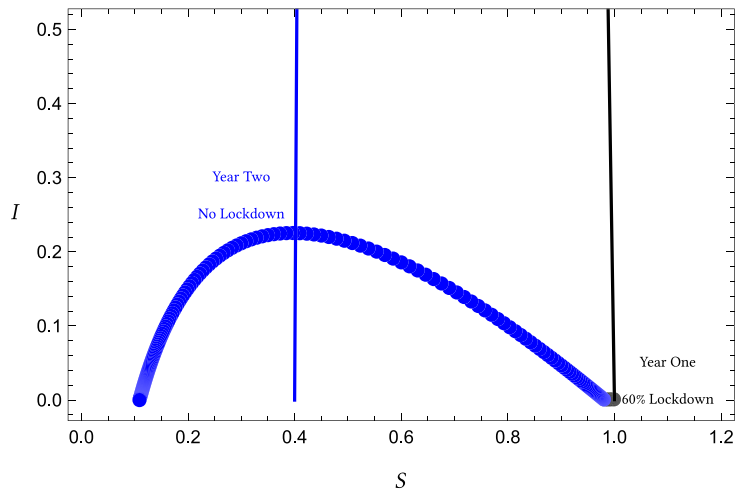


Fig. 3. The Effect of a One Year Lockdown in (S, I) Space.

have $S < 1/\mathcal{R}_0$. That is, assuming $I_0 > 0$, regardless of the initial state, the system converges to a steady state where, at most, $1/\mathcal{R}_0$ agents avoid infection. In particular, as $\mathcal{R}_0 \rightarrow \infty$, the fraction of individuals who never get infected goes to zero (and the fraction of individuals who eventually become infected goes to one). Fig. 2 presents some computed examples of (S_t, I_t) paths starting from $(S_0, I_0) = (0.999, 0.001)$ with only two activities, home and work, and under the assumption that $p_{home} = 0$, $q = 0$, $\gamma = 1/18$, and $M \in \{25\gamma, 16.6667\gamma, 12.5\gamma, 10\gamma\}$, p_{work} is chosen so that the corresponding $\mathcal{R}_0 \in \{2.5, 1.67, 1.25, 1.0\}$. Here, we show the effect of reducing \mathcal{R}_0 by decreasing M from a high of $M = 25\gamma$ (implying $\mathcal{R}_0 = 2.5$, or approximately what epidemiologists consider \mathcal{R}_0 to be without social distancing) to a low of $M = 10\gamma$ (implying $\mathcal{R}_0 = 1$). For the high $M = 25\gamma$, I increases very quickly, and in a short number of periods, almost all people have recovered, or $(S, I) \approx (0.1, 0)$. For the low $M = 10\gamma$, I and S decrease slowly, and the system approaches a steady state with nearly all people having never been infected.

4. Lifting a lockdown

The previous example gives rise to the sobering possibility that a lockdown must go on forever, otherwise, the system returns to one where \mathcal{R}_0 is high and $I_0 > 0$, although $S_0 < 1$. The following example starts with $(S_0, I_0) = (0.999, 0.001)$ and has $M = 10\gamma$ (and thus $\mathcal{R}_0 = 1$) for 365 periods (or one year, assuming a one day period length), then permanently relaxes the lockdown to $M = 25\gamma$ (and $\mathcal{R}_0 = 2.5$). Here, the path after lifting the lockdown is essentially the same as if the lockdown had never been enforced. Fig. 3 shows the result of this exercise in (S, I) space, and Fig. 4 graphs the same exercise showing the infection rate, I_t , over time.

5. Calibration

We assume that the utility function U is the log function. We parameterize the testing cost function as $c(T)^{1+\nu}/(1+\nu)$. We assume that the economy has two activities: work and home. This formulation allows us to relate our results to those in the literature, particularly Eichenbaum et al. (2020b) and Glover et al. (2020).

We set the parameter values to be very similar to equivalent parameter values in Eichenbaum et al. (2020b) and Glover et al. (2020). The parameter values are reported in Table 1. We set the time period to be one week and set the discount factor β , assuming that the vaccine is expected to arrive in 18 months and the annualized discount rate for the planner is 4%. We set the exit rate γ , assuming that the expected length of infection is 18 days. We set the exogenous rate at which unknown infected people become known, $\tilde{\tau}$, using the observation that infected agents are asymptomatic for the first five days of the infection and roughly half of all infected agents never display symptoms (see Glover et al. (2020)). We set the productivity of infected people ξ , following Eichenbaum et al. (2020b). The mortality rate δ is set at 0.5%; see Ferguson et al. (2020). We set the elasticity of substitution at 2 and normalize the productivity in the work sector to be 1 and that in the home sector be 0.1. We set the curvature parameter on the testing cost function, ν , at 1 and choose the parameter c so that the marginal cost of testing is \$50 if 1% of the population is tested. We follow Glover et al. (2020) in setting the cost of treating the infected, Z , so that it is \$7500 over the course of infection. We set the cost of death parameter, D , so that the value of a life is equal to the present discounted value of 15 years of consumption. We set $\pi_{home} = 0.01$ and set π_{work} so that the reproduction rate \mathcal{R}_0 without any policy has an average value of approximately 3 in the first four weeks.

Since the main focus of our analysis is the role of testing and isolation, we experimented with a number of values for the signal probabilities, θ_S and θ_I . For our baseline calibration, we set $\theta_I = 0.38$ and $\theta_S = 0.0044$. To arrive at these numbers, we start with the view that South Korea was particularly effective at pursuing an aggressive test, trace, and isolate policy. We

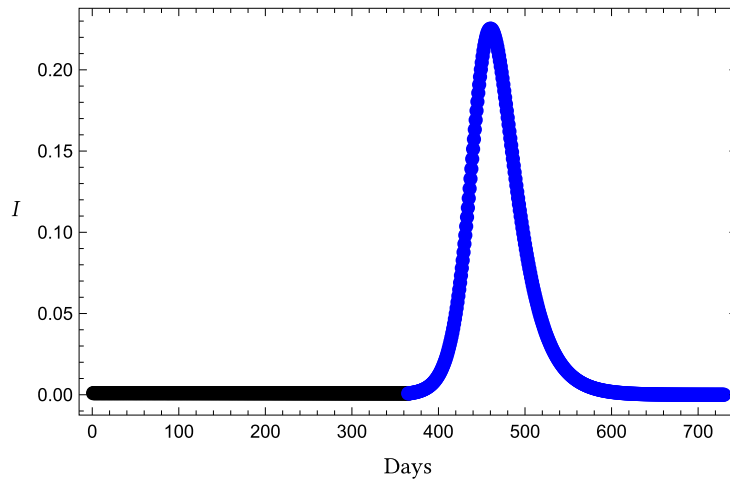


Fig. 4. The Effect of a One Year Lockdown over Time.

Table 1

Model parameters.

β	Discount rate	Weekly model. Vaccine arrival 18 months	0.99
γ	Exit rate	18 day infection period	7/18
$\bar{\tau}$	Prob. of becoming symptomatic	5 day incubation, 1/2 asymptomatic	$\frac{13}{18} \times \frac{1}{2}$
b_1	Productivity of work		1
b_2	Productivity of home		0.1
ξ	Infected productivity loss	Eichenbaum et al. (2020b)	0.8
δ	Death rate	Ferguson et al. (2020)	0.05%
σ	CES parameter		2
ν	Testing cost parameter		1
c	Testing cost parameter	MC of testing \$50 when 1% tested	4.17
Z	Treatment cost	Glover et al. (2020)	6.25
D	Death cost	Value of life 2.8 million dollars	1.84
π_1	Infection rate at home		0.01
π_2	Infection rate at work	\mathcal{R}_0 with no policy 2.5	1.5
θ_1	Signal prob. infected	$\theta_1/\theta_S = 86$ and 1.2% tested (SK)	0.38
θ_S	Signal prob. susceptible	$\theta_1/\theta_S = 86$ and 1.2% tested (SK)	0.0044

think of South Korea as having a state of the art testing and tracking technology. In South Korea, 1.8% of the tests returned a positive result and the proportion of the population that was infected was 0.021%. Viewed through the lens of our model, these data imply

$$\frac{\theta_1 I}{\theta_1 I + \theta_S S} = 0.018.$$

$$\frac{I}{I + S} = 0.0021.$$

Using the observation that early in the pandemic $I + S = 1$, we obtain a value of $\theta_1/\theta_S \approx 86$. The next step in our calibration is to determine what fraction emits the signal. To determine the fraction of population that emits the signal, we use South Korean data in which 1.2% of the population was tested. We assume that in South Korea, everyone who emitted the signal was tested. Thus, we assume that

$$\theta_1 I + \theta_S S = 0.012.$$

We can now determine θ_1 and θ_S once we have a value of I_0 , the initial fraction of the population that was infected. Conceptually we think of a country like the United States having 2% of the population infected in the early phase of the pandemic. We then ask what outcomes would be like under various policies if the US used the South Korean testing and tracking technology, in the sense of the same values for θ_1/θ_S and the same fraction of the population emitting the signal as in South Korea. Clearly, if 2% of the population were infected as opposed to 0.021%, we would expect a larger fraction of the population to emit the signal. Our assumption that only 1.2% of population emits the signal is intended to give us a conservative estimate of the value of testing and isolation. We refer to these parameter values as those associated with the *South Korea-like* technology.

Table 2
Outcomes with and without South Korea-like technology.

Experiment	Welfare gain relative to no intervention	Cumulative deaths	Output loss
No intervention	0	0.48%	1.3%
Opt policy: no testing	0.59%	0.35%	1.94%
Opt policy: untargeted testing	0.71%	0.3%	2.06%
Targ Test ($\theta_S = .0044, \theta_I = .38$)	3.07%	0.15%	1.28%
Targ Isolation ($\theta_S = .0044, \theta_I = .38$)	2.12%	0.26%	1.66%

Using a procedure similar to that for South Korea, we also estimated values for θ_I and θ_S for Germany and Australia. Specifically, for each country, we determined the fraction of all tests that were positive from the start of the epidemic to August 3 (3.02% for Germany and 0.47% for Australia) and the proportion of the population that was infected at the peak of the epidemic for each country (0.088% on April 6 for Germany and 0.029% on August 2 for Australia). We also computed the proportion tested from the start of the epidemic to August 3 (0.48% for Germany and 0.88% for Australia).⁷ Using these numbers and setting the proportion infected at 2% for our baseline country, we obtained $\theta_I = 0.1$ and $\theta_S = 0.003$ for Germany and $\theta_I = 0.1$ and $\theta_S = 0.007$ for Australia. We refer to these parameter values as those associated with the *Germany-like* and *Australia-like* technologies, respectively.

6. Findings

Here, we report on the findings from our quantitative model. Our measure of welfare is the standard compensating variation in consumption widely used in the macroeconomics literature. Specifically, we ask what permanent percentage increase in consumption relative to the no-intervention economy would give the planner the same utility as under our experiments. We report all welfare calculations relative to the no-intervention economy. For reference, we note that the loss in welfare in the no-intervention economy relative to the no-pandemic economy is 6.66%. The welfare changes in our experiments relative to the no-policy case arise from changes in the time paths of output, testing costs, infection costs, and death costs.

We compute welfare measures as well as a partial decomposition of the change in welfare induced by changes in output and death costs. To measure the change in welfare induced by the changes in output, we compute the annuity value of the present discounted value of output in our experiments. To measure the changes in welfare induced by death, we report the cumulative fraction of the population that dies at the end of 52 weeks. We also compute the time paths of the fraction of the population infected I_t , the fraction of population susceptible S_t , cumulative deaths, the reproduction rate \mathcal{R}_{0t} , the fraction of the population infected but not known to be so \tilde{I}_t , consumption, the mass tested, and the marginal cost of testing.

6.1. Benchmark tracking technology

In the first three rows of Table 2, we report outcomes in the no-policy case, optimal policy with no testing, and optimal policy with untargeted testing. We see that optimal policy with no testing yields a welfare gain of roughly 0.6%, and an optimal policy with untargeted testing yields a welfare gain of 0.7% relative to no intervention. That is, untargeted testing delivers very modest welfare gains.

In Fig. 5, we report the time path of outcomes. We see that in all three cases, the economy goes through a severe recession that lasts about three months. We see that optimal policy reduces output at its trough by roughly 40% and by about 20% with no policy intervention. The primary gains to welfare relative to no policy come from a sharp reduction in the cumulative number of deaths. With no policy intervention, roughly 0.5% of the population dies, while with optimal policy, roughly 0.35% of the population dies. In all three experiments, the population eventually reaches herd immunity, though the steady state fractions are very different with and without optimal policy. The figure also shows that the main mechanism by which optimal policy reduces the cumulative death rate is by inducing a sharper recession, which in turn reduces the effective reproduction rate \mathcal{R}_{0t} below 1 and thereby induces a reduction in the proportion of the population that is infected. The findings regarding optimal policy with no testing are broadly similar to the findings in Eichenbaum et al. (2020b) and Glover et al. (2020). Fig. 5 also shows that with untargeted testing, as much as 15% of the population is tested. Note that at its peak, the marginal cost of testing is roughly \$800. In this sense, a relatively small fraction of aggregate resources is allocated to testing. The main reason is that untargeted testing is not very valuable.

Next, we compare outcomes under optimal policy with no testing and untargeted testing with optimal policy when the South Korea-like technology is available. Table 2 shows that the welfare gains to targeted testing when the South Korea-like technology is available are substantial. In particular, welfare rises by 2.5% relative to optimal policy with no testing and 3% relative to no policy. This table also shows that the cumulative deaths are about 0.15% lower once we allow for

⁷ These values, as well as those for South Korea, are obtained from Worldometers (2020) and JohnsHopkins (2020).

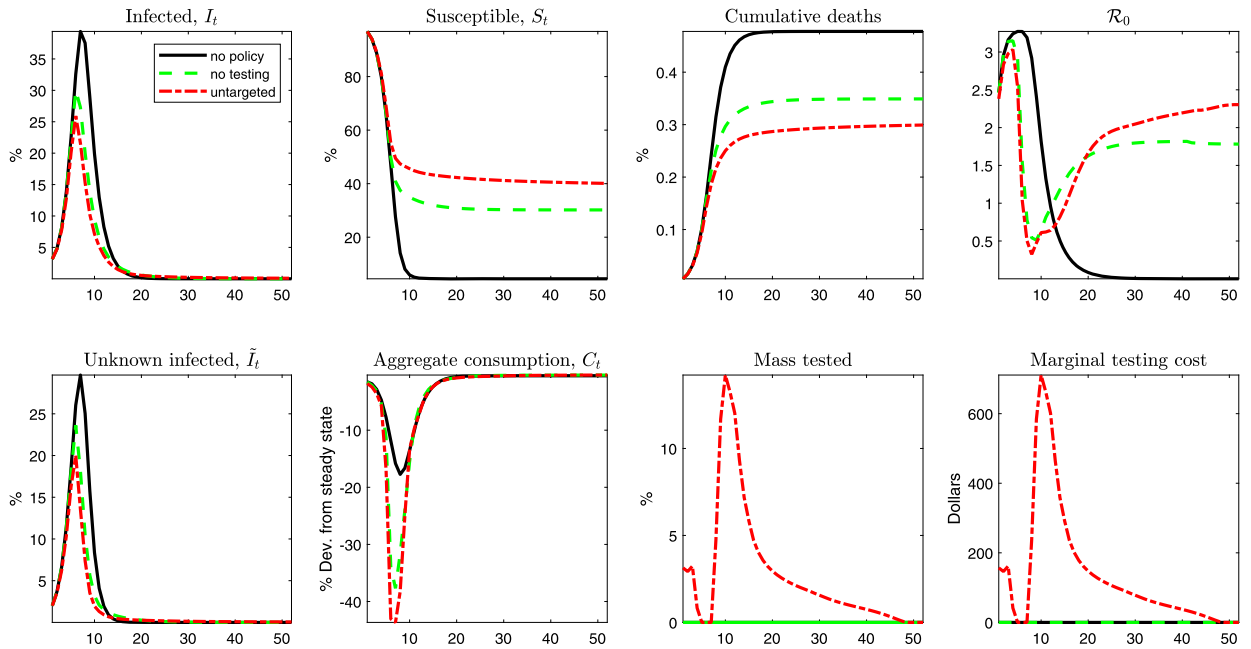


Fig. 5. Time Paths for No Policy, No Testing, and Untargeted Testing.

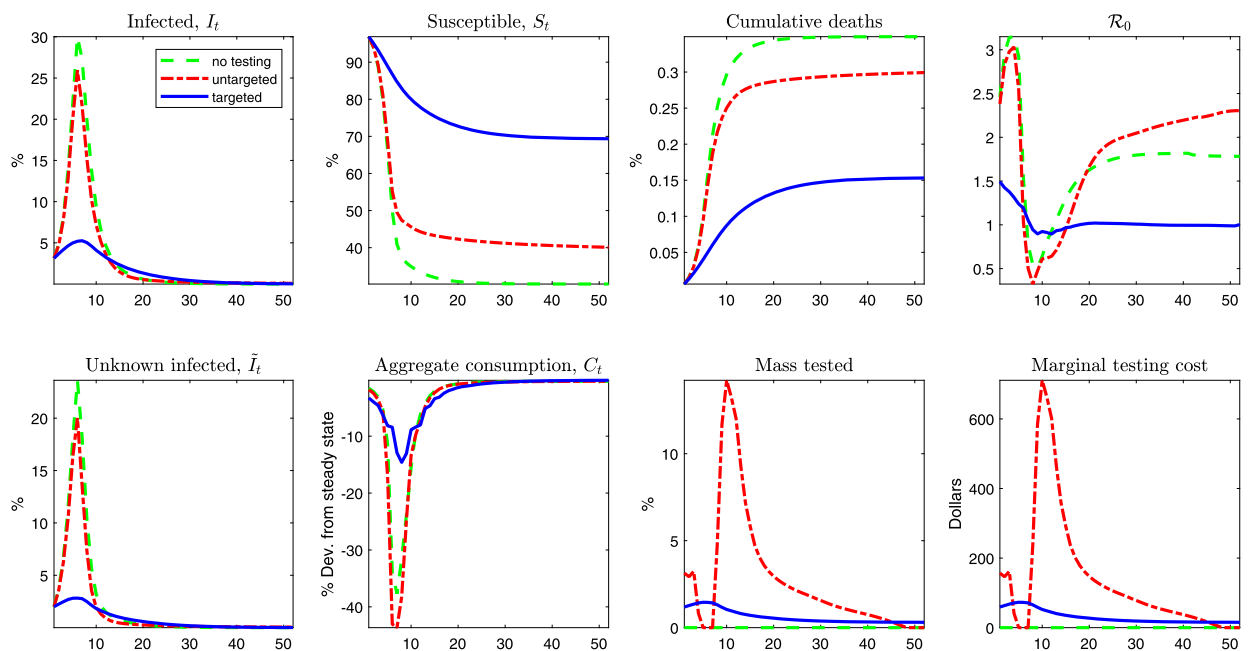


Fig. 6. Time Paths for No Testing, Untargeted Testing, and Targeted Testing (SK).

targeted testing and that the output loss is moderated by about 0.8% relative to no testing. Fig. 6 shows that welfare rises dramatically mainly because a targeted testing policy with South Korea-like technology ensures an initial decline in the effective reproduction rate \mathcal{R}_{0t} and then keeps that rate at around 1. This way of controlling the reproduction rate ensures that the cumulative death rate is substantially lower. Fig. 6 shows that even at the peak of targeted testing, only about 1.5% of the population is tested. The marginal cost of testing, even at its peak, is only about \$70. These results show that relative to untargeted testing, targeted testing when a South Korea-like technology is available is both inexpensive and immensely valuable.

Next, we compare outcomes under our targeted testing model with a targeted isolation model assuming that a South Korea-like technology is available (see Fig. 7). From Table 2 we see that targeted isolation alone generates roughly two-thirds

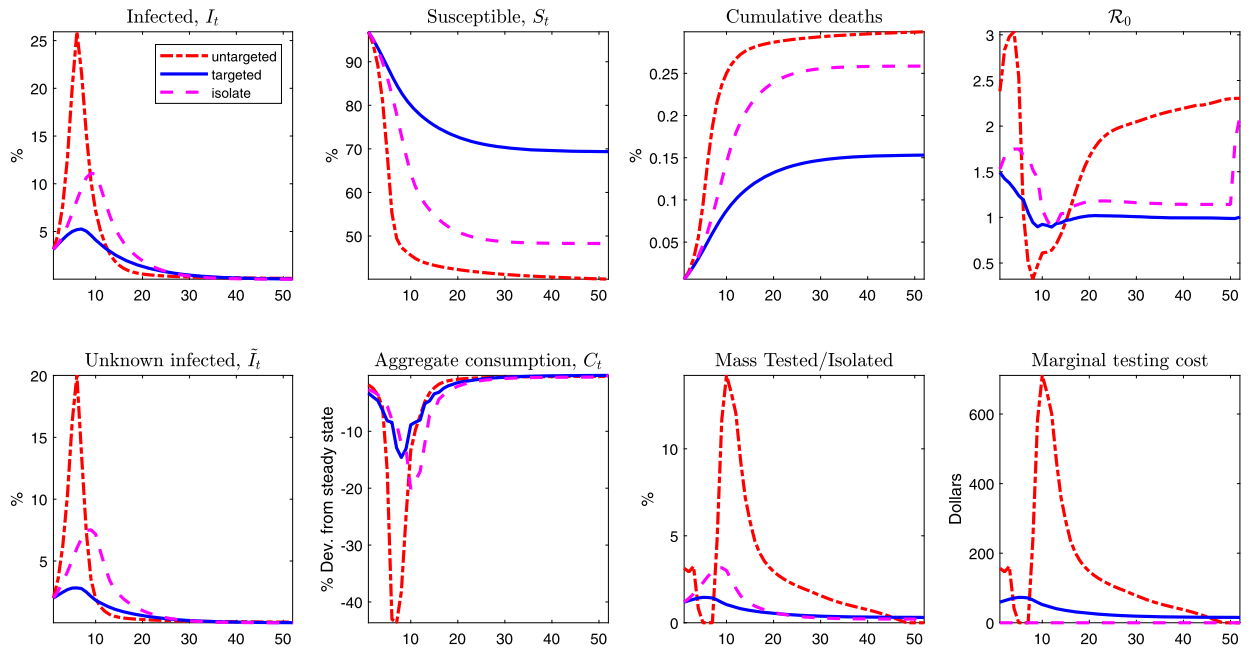


Fig. 7. Time Paths for Untargeted Testing, Targeted Testing (SK), and Isolation (SK).

Table 3
Outcomes with Germany-like technology.

Experiment	Welfare gain relative to no intervention	Cumulative deaths	Output loss
Targ Test ($\theta_s = .003, \theta_l = .1$)	0.69%	0.32%	2.01%
Targ Isolation ($\theta_s = .003, \theta_l = .1$)	0.77%	0.34%	1.93%

Table 4
Sensitivity analysis with respect to signal parameters.

Experiment	Welfare gain relative to no intervention	Cumulative deaths	Output loss
Targeted Test ($\theta_s = .03, \theta_l = .38$)	3.02%	0.15%	1.29%
Targeted Isolation ($\theta_s = .03, \theta_l = .38$)	1.62%	0.30%	1.76%
Targeted Test ($\theta_s = .03, \theta_l = .6$)	5.7%	0.04%	0.19%
Targeted Isolation ($\theta_s = .03, \theta_l = .6$)	3.99%	0.09%	1.56%
Targeted Test ($\theta_s = .03, \theta_l = .15$)	0.78%	0.3%	2.03%
Targeted Isolation ($\theta_s = .03, \theta_l = .15$)	0.71%	0.34%	1.99%

of the welfare gains that come from targeted testing. In this sense, targeted isolation is a very valuable tool if testing is not available.

6.2. Alternative tracking technologies

We consider outcomes that assume Germany-like and Australia-like tracking technologies are available, instead of a South Korea-like one. From Table 3 we see that with a Germany-like technology, targeted testing and targeted isolation generate welfare gains of only 0.69% and 0.77%, respectively. The welfare gains are roughly the same as those with untargeted testing. The results for Australia are very similar and are available on request. Note that the main difference between the South Korea-like and Germany-like technologies is in the parameter θ_l . This parameter measures the fraction of the infected population that emits a signal. In the South Korea-like technology, approximately 40% of the infected people emit the signal, while with the Germany-like technology, only 10% of the infected population emits the signal. We think of this signal as obtained in part from aggressive contact tracing. Our findings with respect to South Korea, Germany, and Australia imply that targeted testing is very valuable if the technology for tracking infected individuals is very good and is not very valuable otherwise. The message of our findings is that the returns to improving the technology for tracking infected individuals can be exceptionally large.

Table 5
Comparison of four and five state models.

Experiment	Welfare gain relative to no-intervention	Cumulative deaths	Output loss
Targ Test ($\theta_s = .0044, \theta_l = .38$, four state)	3.07%	0.15%	1.28%
Targ Test ($\theta_s = .0044, \theta_l = .38$, five state)	2.99%	0.16%	1.25%

This message is reinforced by sensitivity analyses to the values of our signal probabilities. Table 4 shows that if $\theta_l = 0.6$, then welfare rises by about 5.7% relative to the no-policy case. Recalling that the pandemic with no policy delivers a welfare loss of 6.7%, we see that if the technology for tracking infected individuals is sufficiently effective, the welfare loss from the pandemic is only about 1%. For this parameter value, we also see that the cumulative deaths are reduced very significantly to 0.04% from 0.35%. It turns out that the recession that arises from the pandemic is mild. Table 4 also shows that if the technology for tracking infected individuals is much less effective than our benchmark case ($\theta_l = 0.15$), then the welfare, output, and death gains are modest relative to no targeted testing.

6.3. Other sensitivity analyses

Thus far we have considered a four state version of our model. Since the five state model may be more realistic, we computed outcomes for our South Korea-like targeted testing experiment. In Table 5, we report the outcomes for the two models. The table shows that the results for the two models are very similar. The welfare gains to testing are slightly smaller here. The reason is that in the five state model both susceptible people and people who have recovered without having previously been known to be infected emit the signal with probability θ_s while in the four state model only susceptible people emit the signal. Since the signal is less precise in the five state model than in the four state model the planner incurs additional testing costs. In any event the similarity of the two results leads us to be confident that the results will be similar for the other experiments as well. Finally, we conduct sensitivity analyses (available upon request) on the initial fraction of infected agents I_0 . We find the relative gains of testing and isolation policies continue to be substantial. We view this finding as suggesting that even if the pandemic is well under way and testing and isolation policies have not so far been conducted, it is not too late to implement such policies.

7. Conclusion

We have argued that testing and isolation policies can deliver substantial welfare gains in the presence of pandemics. These welfare gains come from a reduced number of cumulative deaths and a shallower recession. Our model can readily be extended to allow for exogenous inflows of agents, some of whom may be infected. Such an extension is useful because we think of our model as one of a particular region or state, rather than of the world. In this context, inter-regional and international migration then introduce new sources of infections. In many situations, the new entering agents can be identified, and the testing and isolation of these agents is clearly valuable. Our findings also suggest that even if a pandemic is well under way, testing and isolation policies are very valuable.

Appendix A. Proof of Proposition 1

Consider the recursive formulation of the programming problem in the model with isolation:

$$V(\tilde{I}, S, R) = \max_{\lambda_i \mu_i, \tau} \left\{ \log Y - Z(I - \tilde{I}) - D(\gamma \delta I) + \beta V(\tilde{I}', S', R') \right\}$$

subject to

$$\tilde{I}' = (1 - \tilde{\tau})(1 - \gamma)\tilde{I} + (1 - \tau\theta_s)S \sum_i \left(\lambda_i \pi_i \frac{\lambda_i (1 - \tau\theta_l) \tilde{I}}{L_i} \right) \quad (19)$$

$$S' = \left[1 - (1 - \tau\theta_s) \sum_i \left(\lambda_i \pi_i \frac{\lambda_i (1 - \tau\theta_l) \tilde{I}}{L_i} \right) \right] S, \quad (20)$$

where R' is given by (8), L_i is given by (14), Y is given by the CES aggregator over y_i given by (15), and $I = 1 - S - R$.

Clearly, the continuation value is decreasing in \tilde{I}' . We will use this result in the proof. Let $\lambda = \lambda_{work}$. Suppose, by way of contradiction, that $\lambda < 1$ and $\tau < 1$. We will construct a variation that increases λ and τ while keeping current output constant. We will show that this variation reduces \tilde{I}' and thereby raises welfare. This contradiction establishes the proof. To this end, totally differentiate output with respect to τ and λ , holding all other variables fixed. We have that

$$d\lambda = d\tau \frac{\lambda \left[S\theta_s + \theta_I \tilde{I} \right]}{\left[S(1 - \tau\theta_s) + [1 - \tau\theta_I] \tilde{I} \right]}. \quad (21)$$

Note that since $\xi = 1$, output equals L . Since output is held constant, the rate of infections is now determined solely by

$$(1 - \tau\theta_s)\lambda^2(1 - \tau\theta_I).$$

Differentiating this expression, substituting from (21), and simplifying, we get that the sign of the rate of infections is given by the sign of

$$\frac{\lambda^2 d\tau}{\left[S(1 - \tau\theta_s) + [1 - \tau\theta_I] \tilde{I} \right]} (1 - \tau\theta_s)(1 - \tau\theta_s)[\theta_I - \theta_s] \left[-S + \tilde{I} \right],$$

which is negative. Q.E.D.

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